

Estimation of Compensation of Employee Function: Panel Data Analysis of National Accounts of Nigeria

Nwabueze Joy Chioma¹, Okoroafor Ikechukwu Benedict^{2*}, Okoroafor-Iyke Florence Eleda³

¹Department of Statistics, Abia State University, Abia State, Nigeria ²Pre-ND Program, Federal College of Agriculture, Ishiagu, Ebonyi State, Nigeria ³ASUBEB. Umuahia, Abia State, Nigeria ²*ekwus2002@vahoo.com

Abstract

Fixed effect models otherwise known as least squares dummy variable regression model was applied to the study of the compensation of employee function. The estimates were compared with ordinary least squares (OLS) estimator. The evaluation criteria for the estimated models are; R-square, F-ratio and restricted F-test. The results showed that the coefficients of the explanatory variables, excluding indirect taxes positively explained compensation of employee. It also revealed that the coefficients of the explanatory variables excluding consumption of fixed capital play significant role in determining what should be paid as compensation to employee who suffer injuries or loses. The restricted F-test for the estimators showed that compensation function has not changed much across the sectors but that it has changed much more over the years.

Key words: Panel data; Fixed effects; Least square dummy variable.

Introduction

The study of two dimensional Panel data is usually of interest in Statistics, Epidemiology and Econometrics. It contains observations on multiple phenomena observed over multiple time periods for the same individuals or households. Data of this type are usually analyzed by panel data regression models. The basic idea underlining the application of panel data regression model is to determine or estimate the functional parameters called the intercept and the slope coefficients as well as to capture the patterns of the relationship for the dependent variable over time series and cross sectional variables. Panel data regression model was developed for situations where the error term, μ_{it} is assumed to vary non-stochastically over i and t, making the model; fixed effects model (FEM) in one dimension. Where the error term, μ_{it} is assumed to vary stochastically over i and t, the model is random effects model (REM), Baltagi (2008). Fixed effects model (FEM) have been applied to both social and Economic problems. Some authors have used fixed effects model in their studies. Notable among them is Ahmed & Sobhi (2009), who studied a comparative study for estimation parameters, used Panel data regression model and their result shows that classical pooling estimator performed well in a fixed effect model only. Cheng and Kamil (1997) showed that fixed effects model may be applied in the analysis of liquidity constraints and firm investment. Treisman (2000) showed that fixed effects model could be extended in various areas, for example to identify the endogeneity of income and potential the instrument. Hsiao & Kamil (1997) showed that fixed effects model was used in selecting the final specification and evaluated the importance of financial constraints on firm's investment decision Lee & Russell (2004) also showed that fixed and random effects models could be applied in Panel data analysis of factors affecting as built roughness of asphalt concrete pavement. In this work, panel data regression model; fixed effects model was studied on "National Accounts of Nigeria". Measurements on cost components of value added at current prices in millions, for four activity sector; Electricity, Telecommunications, Transportation and Education were collected over twenty-six (26) year time period from 1981-2006. The components considered are; Compensation of Employee (CE), Consumption of Fixed Capital (CFC), Indirect Taxes (IT), Value Added at Current Basic Price



(VACBP) and Operating Surplus (OS). These independent variables were actual life data published by National Bureau of Statistics (2007) Nigeria, to measure compensation as referenced.

Table 1. Ordinary least square (OLS) results							
. reg CE CFC IT VACBP OS							
Source	SS	df	df MS		Number of obs = 10 F(4, 99) = 16.1		
Model Residual	2.7447e+10 4.2106e+10	4 99	6.8617e+09 425313673		Prob > F R-squared	= 0.0000 = 0.3946	
Total	6.9553 e+ 10	103	675272094		Adj R-squared = 0. Root MSE = 2		
Œ	Coef.	Std. E	rr. t	P> t	[95% Conf.	Interval]	
CFC	.0528031	.48985	83 0.11	0.914	9191821	1.024788	
П	3764571	.10091	02 -3.73	0.000	5766848	1762294	
VACBP	.3835432	.0997	03 3.85	0.000	.1857108	.5813755	
OS	.9283589	.26813		0.001	.3963186	1.460399	
_cons	1843.832	2305.7	08 0.80	0.426	-2731.193	6418.856	

Table 2. Results of fixed effects when the error term is fixed

. xtreg CE CFC IT VACBP OS,fe

Fixed-effects (within) regression Group variable: id					f obs f group	= 0S =	104 4
	= 0.3702 n = 0.7915 l = 0.3933			Obs per	group:	min = avg = max =	26 26.0 26
corr(u_i, Xb)	= 0.0594			F(4,96) Prob > F		=	14.11 0.0000
CE	Coef.	Std. Err.	t	P> t	[95%	Conf.	Interval]
CFC IT VACBP OS _CONS	.2162725 3753087 .3820496 .8046859 1856.271	.5122012 .10352 .1023537 .3073387 2344.345	0.42 -3.63 3.73 2.62 0.79	0.674 0.000 0.000 0.010 0.430	8004 5807 .1788 .1940 -2797	7942 8791 5234	1.232984 1698231 .58522 1.414748 6509.758
sigma_u sigma_e rho	3282.8113 20757.125 .02440216	(fraction	of variar	nce due to	u_i)		
F test that a	ll u_i=0:	F(3, 96) =	0.58	58 Prob > F = 0.6326			

Theoretical Framework

The basic theoretical problem which will be considered in this work is based on the application of ordinary least squares (OLS) and fixed effects

> model, using least square dummy variable regression model. A Panel data model as given by Gujarati (2006) is:

 $Y_{it} = \beta_1 + \beta_2 X_{2it} + \beta_3 X_{3it} + \ldots + \beta_k X_{kit} + \mu_{it}$ (1)

i = 1, 2, ..., k and t = 1, 2, 3, ..., n

where i stands for the ith cross sectional units (sectors)and t stands for the tth time period (years) which are the cross sectional identifier and the time identifier respectively. β_1 is s constant. β_2 , β_3 , ..., β_k are the regression coefficients. Y_{it} is the dependent variable while X₂, X₃, ..., X_k are the independent variables.

Analytical Framework Ordinary least squares (OLS)

The ordinary least squares (OLS) regression applied to investigate Compensation of Employee function when coefficients are constant across time and individual sectors is achieved by pooling all 104 the observations together, Thus, 4 equation (1) could be written as:

 $CE_{it} = \beta_1 + \beta_2 CFC_{2it} + \beta_3 IT_{3it} + \beta_4 VACBP_{4it} + \beta_5 OS_{5it} + \mu_{it}$

i = 1, 2, ..., 4 and t = 1, 2, 3,, 26

where the variables are as defined in the introduction. Stacking all the twenty-six observations for each sector units, one on top of the other, thus, given in all, 104 observations for each of the variables in the model, disregarding the space and time dimensions of the pooled data and estimate the usual ordinary least squares regression. The results obtained are



shown in Table 1.

Fixed effects model (FEM)

In fixed effects model, the error term μ_{it} is assumed to vary non- stochastically over i and t making the fixed effects model in one dimension. In fixed effects, we impose time independent effect on each entity that are possibly correlated with the regressors, Gujarati (2006); Gujarati & Porter (2009). Fixed effects model used in this work is given as:

 $CE_{it} = \beta_{1i} + \beta_2 CFC_{2it} + \beta_3 IT_{3it} + \beta_4 VACBP_{4it} + \beta_5 OS_{5it} + \mu_{it} (2)$

where $\mu_{it} = \mu_i + v_{it}$ and μ_i are individual sector specific time-invariant effect. The subscript i on the constant term suggests that the constants of the four sector may be different which may be as a result of some special features of each of the sectors such as management style and policies. The results obtained are presented in Table 2.

Least squares dummy variables (LSDV) regression models

Least squares dummy variable is a method which takes into accounts the individuality of each sector. This is achieved by allowing the constant to vary for each sector but still assume that the regression coefficients are constant across sector and/ or time periods, Hsiao (2003). Assuming that the constant varies across individual sectors but the regression coefficients are constant. Equation (2) becomes:

 $\begin{array}{rcl} CE_{it} &=& \alpha_1 + & \alpha_2 D_{2i} + & \alpha_3 D_{3i} + & \alpha_4 D_{4i} + & \beta_2 CFC_{2it} & + \\ \beta_3 IT_{3it} + & \beta_4 VACBP_{4it} + & \beta_5 OS_{5it} + & \mu_{it} & (3) \end{array}$

where $D_2 = 1$ if the observation belongs to Telecommunications, $D_2 = 0$ if the observation does not belong to Telecommunications;

 $D_3 = 1$ if the observation belongs to Transportation, $D_3 = 0$ if the observation does not belong to Transportation services;

 $D_4 = 1$ if the observation belongs to Education, $D_4 = 0$ if the observation does not belong to Education. α_1 represents the constant of Electricity while α_2 , α_3 and α_4 are the differential constant coefficients, telling us by how much the intercepts of Telecommunication, Transportation and 154

Education differ from the intercept of Electricity, which is our comparison sector. We used only three dummies to avoid falling into dummy variable trap (the situation of perfect collinearity). The results are presented in Table 3. When we assume that the regression coefficients are constant but the constant coefficient varies across individual time periods. Just as we used dummy variables to account for

 Table 3. Results of Least Squares Dummy Variable for Individual Sector Effects

reg CE CFC IT VACBP OS D2 D3 D4

Source	SS	df	MS		Number of obs F(7, 96)	= 9.35
Model Residual	2.8191e+10 4.1362e+10		272e+09 0858229		Prob > F R-squared Adj R-squared	= 0.0000 = 0.4053 = 0.3619
Total	6.9553e+10	103 67	5272094		Root MSE	= 20757
CE	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
CFC IT VACBP OS D2 D3 D4 CONS	.2162725 3753087 .3820496 .8046859 -7687.846 -1854.13 -3612.011 5144.768	.5122012 .10352 .1023537 .3073387 6351.404 6408.816 7048.79 4727.201	0.42 -3.63 3.73 2.62 -1.21 -0.29 -0.51 1.09	0.674 0.000 0.000 0.229 0.773 0.610 0.279	8004388 5807942 .1788791 .1946234 -20295.28 -14575.53 -17603.75 -4238.652	1.232984 1698231 .58522 1.414748 4919.591 10867.27 10379.73 14528.19

individual sector effects, individual time effects can easily be taken care of by introducing time dummies, one for each year in the sense that compensation function shift over time because of factors such as technological changes, changes in government, taxation policies and external factors as conflicts and war. Thus, equation (2) becomes: $CE_{it} = \lambda_0 + \lambda_1 DUM_1 + \lambda_2 DUM_2 + \dots + \lambda_{25} DUM_{25} + \dots$ $\beta_2 CFC_{2it} + \beta_3 IT_{3it} + \beta_4 VACBP_{4it} + \beta_5 OS_{5it} + \mu_{it}$ (4)where $DUM_1 = 1$ for observation in 1981 and 0 otherwise; $DUM_2 = 1$ for observation in 1982 and 0 otherwise; $DUM_3 = 1$ for observation in 1983 and 0 otherwise; . . .; $DUM_{25} = 1$ for observation in 2005 and 0 otherwise. 2006 was treated as the base year whose intercept value is λ_0 . The differential intercept coefficients λ_1 , λ_2 , λ_3 , ..., λ_{25} tell us by how much the intercepts of 1981, 1982, 1983, ..., 2005 differ from the intercept of 2006 which is our comparison year. The results of model (4) are presented in Table 4.



Table 4. Results of Least Squares Dummy Variable for Individual Time Effects

. reg CE CFC IT VACBP OS DUM1- DUM25

Source	SS	df	df MS			Number of obs F(29, 74)	
Model	5.0315e+10	29	1.7350e+09			P(25, 74) Prob > F	= 0.0000
Residual	1.9238e+10	74	2599	975423		R-squared	= 0.7234
	C 0552. 10	102	(75)	72004		Adj R-squared	
Total	6.9553e+10	103	0/54	272094		Root MSE	= 16124
CE	Coef.	Std.	Err.	t	P> t	[95% Conf.	Interval]
CFC	-2.333204	.4619	527	-5.05	0.000	-3.253665	-1.412743
IT	2632973	.0800	729	-3.29	0.002	4228461	1037485
VACBP	.2671535	.0792	349	3.37	0.001	.1092745	.4250325
0S	.7290806	.2366	246	3.08	0.003	.2575958	1.200565
DUM1	-116810.5	14	918	-7.83	0.000	-146535.2	-87085.68
DUM2	-116856	14923	.19	-7.83	0.000	-146591.1	-87120.93
DUM3	-117530.3	14892		-7.89	0.000	-147205.2	-87855.35
DUM4	-116685.2	1490		-7.83	0.000	-146387.9	-86982.6
DUM5	-116485.7	1488		-7.83	0.000	-146144.3	-86827.07
DUM6	-116440.2	14881		-7.82	0.000	-146092.1	-86788.28
DUM7	-116149.8	14852		-7.82	0.000	-145743.7	-86555.89
DUM8	-115720	14809		-7.81	0.000	-145228.5	-86211.49
DUM9	-115411.1	14773		-7.81	0.000	-144847.1	-85975.04
DUM10	-115030.6	14733		-7.81	0.000	-144388.4	-85672.8
DUM11	-114769.2	1470		-7.80	0.000	-144077	-85461.53
DUM12	-115018.1		695	-7.83	0.000	-144298.5	-85737.64
DUM13	-113982.6	14658		-7.78	0.000	-143190.6	-84774.71
DUM14	-113210.3	1460		-7.75	0.000	-142314.8	-84105.85
DUM15	-111789.9	14527		-7.70	0.000	-140735.8	-82843.99
DUM16	-110582.5	14457		-7.65	0.000	-139389.6	-81775.38
DUM17	-109836.5	1440		-7.62	0.000	-138546.4	-81126.54
DUM18	-109047.3	14355		-7.60	0.000	-137650.6	-80444.04
DUM19	-108367.7	14302		-7.58	0.000	-136866.9	-79868.56
DUM20	-107691.4	14249		-7.56	0.000	-136083.5	-79299.27
DUM21	-102418.8	13648		-7.50	0.000	-129614.3	-75223.4
DUM22	-102288.7	13384		-7.64	0.000	-128958.6	-75618.73
DUM23	-100417.3	13213		-7.60	0.000	-126746	-74088.66
DUM24	-98476.71	13114		-7.51	0.000	-124608.8	-72344.63
DUM25	-92559.07	12595		-7.35	0.000	-117655.7	-67462.43
_cons	117431.3	12635	.22	9.29	0.000	92255.03	142607.5

Results and Discussion

The results of Tables 1 to 4 show that the coefficients of the regression have positive signs consistent with a priori expectation, except that of Indirect Taxes which has a negative sign. Implying that, Consumption of Fixed Capital (CFC), Value Added at Current Basic Price (VACBP) and Operating Surplus (OS) positively explained Compensation of Employee function. In other words, an increase in any of these variables will increase Compensation of Employee while an increase Indirect Taxes will lead to a reduction of compensation of employee and vice versa. The results revealed that the coefficients of Indirect Taxes (IT), Value Added at Current Basic Price

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(VACBP) and Operating Surplus (OS) are highly statistically significant. Implying that the percentage paid as tax or the amount paid as tax, Current Basic prices at which sectors offer their products for sale and Surplus, a situation where revenue generated is greater the expenditure, play major/significant role in determining what should be paid compensation to staff (workers) who suffer injury or loss. Only the coefficient of Consumption of Fixed Capital (CFC) is not statistically significant. This is because majority of the fixed Capitals (tangible assets) might be old or broken down becoming liabilities to the sectors instead of being productive. The coefficients of determination, R^2 of 0.3946 and 0.3702 in Tables 1 and 2 respectively are relatively low, indicating about 39.46 and 37.02 percent of the total variations observed in the dependent variable actually accounted for by were the independent variables included in the models 1 and 2. Table 3 showed that all the estimated coefficients of the sector units are individually statistically not significant, as p- values 0.229, 0.773 and 0.610 of the estimated t-coefficients are small. The constant values of the four sector units are statistically different; being 5144.768 for Electricity; -2543.08 (5144.768 - 7687.846) for Telecommunications services; 3290.64 (5144.768 1854.13) _ for

Transportation and 1532.76 (5144.768 – 3612.011) for Education. The differences in their constant values may be due to unique features of each sector, such as managerial ability technological changes, changes in government, taxation policies etc. The negative value for Telecommunications services suggests that the differential constant of Telecommunications services differ by -2543.08 from the constant of Electricity. The R² value of 0.4053 shows an increment of 0.0107 over that of Table1. The restricted F-test of 0.58 is not statistically significant; therefore suggesting that the differential intercepts of model (3) are the same. The results in Table 4 revealed that individual time dummy coefficients are statistically



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significant. Again all the time dummy coefficients have negative signs showing that compensation of employee is negatively explained by individual time periods (years). Coefficient of determination, R^2 value of 0.7234 was reasonably high when compared with the R^2 value of 0.3946 in Table 1; which shows an increment of 0.3288. The restricted F-test value of 29.32 shows that the increment is highly statistically significant, this probably suggests that model (4) seems not valid, suggesting that the differential time intercepts for model (4) are not the same.

Conclusion

The regression model estimators used in this work showed that the coefficients of all the explanatory variables except that of indirect taxes are positively related to Compensation of Employee. Also, all the coefficients of the variables explanatory excluding that of Consumption of Fixed Capital are highly statistically significant. However, on the bases of the restricted F-test for the estimated model (3), compensation function has not changed much across sector but restricted F-test for estimated model (4), compensation function has changed much more over the years.

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